

# Contents

- 9.3 Synopsis of discrete spatial interaction models . . . . . 14
  - 9.3.1 Analysis of Markov models . . . . . 14
  - 9.3.2 A matrix formalism . . . . . 17
  - 9.3.3 Gauss Markov Random Field . . . . . 17

## 9.3 Synopsis of discrete spatial interaction models

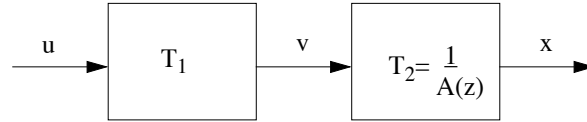


Figure 9.1: Synthesis of the random field.

We restrict ourselves to discrete spatial interaction models with real or integer valued grey level intensities. The random field  $\mathbf{X}$  is *generated* by transformations of a random sequence  $\mathbf{u}$  as samples of a certain stochastic process  $\mathbf{U}$ , which will be specified later. In Fig. 9.1 is presented the diagram for the synthesis of the random field  $\mathbf{X}$ . The stochastic model of  $\mathbf{X}$  is defined by the transformations  $T_1$  and  $T_2$  and the characteristics of  $\mathbf{U}$ .  $T_1$  and  $T_2$  are not limited by any physical realizability considerations as is the case in system theory. Transformation  $T_2$  is defined as

$$T_2 = \frac{1}{A(z)} \quad (9.1)$$

where

$$A(z) = 1 - \sum_{r \in \mathcal{N}_s} \theta_r z^r \quad (9.2)$$

$\theta_r$  are the model parameters, and the following notation is used

$$z^r x_s = x_{s+r} = x_{s_1+r_1, s_2+r_2} \quad (9.3)$$

$r = (r_1, r_2)$  and  $s = (s_1, s_2)$  are sites of the lattice  $\mathcal{L}$ , and  $\mathcal{N}_r$  is the vicinity of site  $r$ . We will also use the alternative definition for the spectral representation

$$A_D(r) = A \left( z = e^{j \frac{2\pi r}{M}} \right) \quad (9.4)$$

where  $M \times M$  is the size of the lattice  $\mathcal{L}$ .

The transformation  $T_1$  has different expressions for different classes of models. In Table 9.1 we give the definition for only three such models.

### 9.3.1 Analysis of Markov models

In this lecture we restrict the discussion to only a particular class of Markov models. The transformation  $T_2$  will be written as:

$$x_s - \sum_{r \in \mathcal{N}_s} \theta_r x_{s+r} = \sqrt{\sigma^2} v_s \quad (9.5)$$

Model	$T_1$
Simultaneous autoregressive (SAR)	$T_1 = 1$
SAR moving average (SARMA)	$T_1 = B(z) = 1 + \sum_{r \in \mathcal{N}_s} \beta_r z^r$
Conditional Markov (MRF)	<p><math>T_1</math> is defined implicitly from the power spectrum</p> $\Phi_v(\omega) = A(e^{j\omega})$ <p>The following notation is used</p> $T_1 = \sqrt{A(z)}$ <p><b>Obs.:</b> If there is an <math>A_1(z)</math> such that</p> $A_1(z) A_1(z) = A(z)$ <p>the model is a SAR model. If <math>A(z)</math> has no factorization, the definition has no meaning. In one dimension a symmetric polynomial <math>A(z)</math> is always factorizable. This is one of the key differences and difficulties of the two-dimensional problems.</p>

Table 9.1: Stochastic models and the associated transformations.

with  $\theta_r = \theta_{-r}$ , the model is defined on a symmetric neighborhood. The process  $\mathbf{v}$  and the transformation  $T_1$  are implicitly defined by the following conditions:

$$\mathcal{E}[v_s|x_r; r \neq s] = 0 \quad (9.6)$$

i.e.  $v_s$  is independent of  $x_r$  for sites  $r \neq s$ . Hence

$$\mathcal{E}[x_s|x_r; r \neq s] = \sum_{r \in \mathcal{N}_s} \theta_r x_{s+r} \quad (9.7)$$

$\mathbf{X}$  is a MRF with respect to the symmetric neighbor set  $\mathcal{N}$  on the lattice  $\mathcal{L}$ .

If the process  $v$  is Gaussian we can write:

$$\begin{aligned} p(x_s|x_r; r \neq s) &= p(x_s|x_{s+r}; r \in \mathcal{N}_s) \\ &= \frac{1}{\sqrt{2\pi} \sigma^2} e^{-\frac{(x_s - \sum_{r \in \mathcal{N}_s} \theta_r x_{s+r})^2}{2\sigma^2}} \end{aligned} \quad (9.8)$$

Further we assume  $X$  to be stationary and evaluate its power spectral density. We will use the Wiener-Khintchine theorem (Eq. 7.31). Eq. 9.5 is multiplied by  $x_{s+s'}$  and the expectation of the product is evaluated. Using the statistical independence in Eq. 9.6 we obtain the autocorrelation function

$$R_X(r') = \sum_{r \in \mathcal{N}} \theta_r R_X(r+r') + \sigma^2 \delta(r') \quad (9.9)$$

where

$$\delta(r') = \begin{cases} 0 & \text{if } r' \neq (0,0) \\ 1 & \text{if } r' = (0,0) \end{cases} \quad (9.10)$$

After evaluation of the Fourier transform of both sides, we obtain:

$$\left(1 - \sum_{r \in \mathcal{N}} \theta_r z_r\right) \Phi_{\mathbf{X}}(\omega) = \sigma^2 \quad (9.11)$$

Thus

$$\Phi_{\mathbf{X}}(\omega) = \frac{\sigma^2}{1 - \sum_{r \in \mathcal{N}} \theta_r \exp\left(j\left(\frac{2\pi}{M}r_1\omega_1 + \frac{2\pi}{M}r_2\omega_2\right)\right)} \quad (9.12)$$

where  $\omega = (\omega_1, \omega_2)$ , with  $0 \leq \omega_1, \omega_2 \leq M-1$ . Using the model definition we observe that

$$\Phi_{\mathbf{X}}(\omega) = \frac{\sigma^2}{\|A \exp(j\omega)\|^2} \Phi_{\mathbf{v}}(\omega) \quad (9.13)$$

which defines the process  $\mathbf{v}$ .

### 9.3.2 A matrix formalism

Until now we wrote all equations for a generic site  $s$  of the lattice  $\mathcal{L}$ . The characterization of the random field, as given in Eq. 9.5, is actually written for all  $M \times M$  sites of the lattice  $\mathcal{L}$ . Further we assume the lattice  $\mathcal{L}$  to have circular symmetry as introduced in section 8.7, and use the stack vector representation  $\mathbf{x}$  of the image  $x_s, s \in \mathcal{L}$ . Thus we can write the model Eq. 9.5 in matrix form:

$$\sqrt{\mathcal{A}(\theta)}\mathbf{x} = \sigma \mathbf{v} \quad (9.14)$$

Where  $\mathbf{v}$  is the stack vector used for the representation of the process  $v$ , and  $\mathcal{A}(\theta)$  is a symmetric block circulant matrix.

$$\mathcal{A}(\theta) = \begin{bmatrix} \mathcal{A}_{1,1} & \mathcal{A}_{1,2} & \cdots & \mathcal{A}_{1,M} \\ \mathcal{A}_{1,M} & \mathcal{A}_{1,1} & \cdots & \mathcal{A}_{1,M-1} \\ \vdots & \vdots & \ddots & \vdots \\ \mathcal{A}_{1,2} & \mathcal{A}_{1,3} & \cdots & \mathcal{A}_{1,1} \end{bmatrix} \quad (9.15)$$

with each  $\mathcal{A}_{1,i}$  a  $M \times M$  circulant matrix,

$$\mathcal{A}_{1,i} = \mathcal{A}_{1,M-i+2}^T \quad (9.16)$$

Now we can write the expression for the joint p.d.f. of the entire image assuming the knowledge of the p.d.f. of the process  $v$ .

$$p(\mathbf{x}) = \frac{1}{\sigma^{\frac{M^2}{2}} |\det \mathcal{A}(\theta)|^{\frac{1}{2}}} \prod_{r \in \mathcal{L}} p(v_r) \quad (9.17)$$

where  $\det \mathcal{A} = \prod_{r \in \mathcal{L}} \|A_D(r)\|^2$ . If  $\mathbf{v}$  is a Gaussian process

$$p(\mathbf{x}) = \frac{\exp\left(-\frac{1}{2\sigma^2} \mathbf{x}^T \mathcal{A}(\theta) \mathbf{x}\right)}{(2\pi\sigma^2)^{\frac{M^2}{2}} \prod_{r \in \mathcal{L}} A_D^{\frac{1}{2}}(r)} \quad (9.18)$$

### 9.3.3 Gauss Markov Random Field

Thus, in summary, a Gauss Markov Random Field (GMRF) defined on the lattice  $\mathcal{L}$  and having symmetric neighborhood is characterized by

$$x_s = \sum_{r \in \mathcal{N}_s} \theta_r x_{s+r} + \sqrt{\sigma^2} v_s \quad (9.19)$$

$$\sqrt{\mathcal{A}(\theta)}\mathbf{x} = \sigma \mathbf{v} \quad (9.20)$$

With  $\mathbf{v}$  a zero mean, unit variance Gaussian noise having

$$\mathcal{E}[v_r \cdot v_{r+s}] = \begin{cases} 1 & \text{if } r = (0,0) \\ -\theta_r & \text{if } r \in \mathcal{N}_s \\ 0 & \text{else} \end{cases} \quad (9.21)$$

and  $\theta_r = \theta_{-r}$ . The covariance matrix  $C_X$  of  $X$  is positive definite. The joint p.d.f. of  $X$  is

$$p(\mathbf{x}) = \frac{1}{\sigma^{\frac{M^2}{2}} |\det C_X|^{\frac{1}{2}}} e^{-\frac{1}{2\sigma^2} \mathbf{x}^T C_X^{-1} \mathbf{x}} \quad (9.22)$$

The Markov property

$$\begin{aligned}
 p(x_s|x_r; r \neq s) &= p(x_s|x_{s+r}; r \in \mathcal{N}_s) \\
 &= \frac{1}{\sqrt{2\pi} \sigma^2} e^{-\frac{(x_s - \sum_{r \in \mathcal{N}_s} \theta_r x_{s+r})^2}{2\sigma^2}}
 \end{aligned} \tag{9.23}$$

and the power spectrum is

$$\Phi_{\mathbf{X}}(\omega) = \frac{\sigma^2}{1 - \sum_{r \in \mathcal{N}} \theta_r \exp\left(j\left(\frac{2\pi}{M} r_1 \omega_1 + \frac{2\pi}{M} r_2 \omega_2\right)\right)} \tag{9.24}$$

A GMRF is a special case of GRF. A Gaussian process is a special case of GMRF whose Gibbs energy consists only of single site clique potentials. It has no contextual interaction.

# Bibliography

- [1] G. Winkler, *Image Analysis, Random Fields and Dynamic Monte Carlo Methods*, Springer, 1995
- [2] S. Li, *Markov Random Field Modeling in Computer Vision*, Springer, 1995